# Machine Learning FIB, Master in Innovation and Research in Informatics

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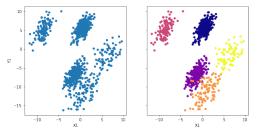
Topic 2: Probabilistic clustering

## Outline

- 1. What is clustering
- 2. K-means clustering
- 3. Mixture of Gaussians (MoG)
- $4.\ \,$  Expectation Maximization for learning MoG

## What is clustering

The goal of **clustering** is to partition a data sample into groups ("clusters") in such a way that observations in the same cluster tend to be more similar than observations in different clusters



Input data is embedded in a d-dimensional space with a similarity/dissimilarity function defined among elements in the space, which should capture relatedness among elements in this space :

- lacksquare elements e
- ▶ a "cluster" is a compact group that is separated from other groups or elements outside the cluster

## What is clustering

There is no consensus<sup>1</sup> on how to measure in a concise way (mathematically) these ideas, and different algorithms capture them in their own way. Thus, there is a large variety of clustering algorithms:

- ▶ Hierarchical bottom-up/top-down: single or average linkage, Ward, . . .
- ▶ Probabilistic use MoGs (e.g. k-means and E-M)
- ▶ Possibilistic use fuzzy set memberships (e.g. fuzzy c-means)
- Algorithmic greedy/hill-climbing (swapping elements between clusters, e.g. PAM)
- ▶ Spectral use the spectrum (eigenvalues/vectors) of the data similarity matrix to perform dimensionality reduction before clustering in fewer dimensions
- ▶ Density-based finds connected dense regions in the data space (e.g. DBSCAN)

 $<sup>^1{\</sup>rm Kleinberg's}$  paper An Impossibility Theorem for Clustering argues that no consensus can exist.

# Clustering is hard

▶ The number of ways one can partition a set of *N* elements into *K* groups is astronomical, it is given by the Stirling number of the second kind:

$$S(N,K) = \frac{1}{K!} \sum_{i=1}^{K} (-1)^{i} {K \choose i} (K-i)^{N}$$

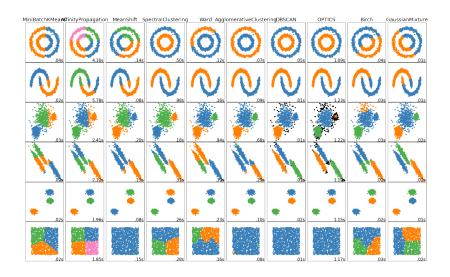
▶ Adding over all possible K = 1, ..., N, we get that the number of ways to partition a set of N elements is given by the N-th Bell number:

$$B(N) = \sum_{K=1}^{N} S(N, K)$$

To get a sense of these numbers:

$$S(10,4) = 35.105$$
  $S(19,4) \approx 10^{10}$   $B(71) \approx 4 \times 10^{71}$ 

# Many algorithms for many situations..



# In this lecture, we'll cover

- ▶ K-means clustering
- ▶ Expectation-Maximization for Gaussian mixture models



#### Input:

- ▶ a dataset  $D = \{\mathbf{x}_1, \dots, \mathbf{x}_n\}$  with  $\mathbf{x}_i \in \mathbb{R}^d$ , and
- $\triangleright$  an integer K > 1 (a hyper-parameter) denoting the number of desired clusters

#### Main intuition:

- each cluster represented by a cluster center  $\mu_k \in \mathbb{R}^d$  for k = 1, ..., K
- ▶ each point should be closest to its assigned cluster center

In a "good cluster" all of its points should be close to its assigned cluster center, and so we would expect

$$\sum_{i:\mathbf{x}_i \text{ is assigned} \atop \text{to cluster } k} \|\mathbf{x}_i - \mu_k\|^2$$

to be small, where  $\|\mathbf{x}_i - \mu_k\|^2$  is the Euclidean distance between  $\mathbf{x}_i$  and cluster center  $\mu_k$ .

Cost function to optimize

We introduce a set of indicator variables:

$$r_{ik} = \begin{cases} 1 & \text{if } \mathbf{x}_i \text{ is assigned to cluster } k \\ 0 & \text{otherwise} \end{cases}$$

And an objective (cost) function:

$$\mathcal{J}(\mu, \mathbf{r}) = \sum_{k=1}^{K} \sum_{i=1}^{n} r_{ik} \|\mathbf{x}_{i} - \mu_{k}\|^{2}$$

The goal is to find cluter centers  $\{\mu_k\}_k$  and assignments  $\{r_{ik}\}_{ik}$  that minimize

$$\mathcal{J}(\mu, \mathbf{r}) = \sum_{k=1}^{K} \sum_{i=1}^{n} r_{ik} \|\mathbf{x}_{i} - \mu_{k}\|^{2}.$$

Unfortunately, this is an **NP-hard problem**. So we will use an alternative procedure to try to minimize it. It is only guaranteed to find **local minima**.

It is based on the fact that:

- a) For fixed cluster centers  $\mu_k$ , it is easy to optimize cluster assignments  $r_{ik}$
- b) For fixed cluster assignments  $r_{ik}$ , it is easy to optimize cluster centers  $\mu_k$

Optimizing the cost function  $\mathcal{J}$  for fixed  $\mu_k$ 

Assume fixed cluster centers  $\mu_k$  for  $k = 1, \ldots, K$ .

The optimal way to assign points to clusters is by assigning each point to its nearest cluster center:

$$r_{ik} := \begin{cases} 1 & \text{if } k = \arg\min_{k'} \|\mathbf{x}_i - \mu_{k'}\|^2 \\ 0 & \text{otherwise} \end{cases}$$

To see this, it is enough to observe that if a point  $\mathbf{x}_i$  is assigned to a cluster center  $\mu_j$  instead of being assigned to its closest  $\mu_k$ , then the cost function can be improved because its contribution to the cost is given by  $\|\mathbf{x}_i - \mu_j\|^2 > \|\mathbf{x}_i - \mu_k\|^2$ 

Optimizing the cost function  $\mathcal{J}$  for fixed  $r_{ik}$ 

Assume fixed  $r_{ik}$  for i = 1, ..., n and k = 1, ..., K.

Following usual procedure, we differentiate and equate to 0 in order to find the minimum.

$$\frac{\partial}{\partial \mu_j} \mathcal{J}(\mu_1, \dots, \mu_K) = \sum_{k=1}^K \sum_{i=1}^n \frac{\partial}{\partial \mu_j} r_{ik} \|\mathbf{x}_i - \mu_k\|^2$$

$$= \sum_{i=1}^n r_{ij} \frac{\partial}{\partial \mu_j} (\mathbf{x}_i - \mu_j)^T (\mathbf{x}_i - \mu_j)$$

$$= \sum_{i=1}^n r_{ij} \frac{\partial}{\partial \mu_j} \left\{ \mathbf{x}_i^T \mathbf{x}_i - 2\mu_j^T \mathbf{x}_i + \mu_j^T \mu_j \right\}$$

$$= \sum_{i=1}^n r_{ij} \left\{ -2\mathbf{x}_i + 2\mu_j \right\}$$

$$= -2 \sum_{i=1}^n r_{ij} \mathbf{x}_i + 2\mu_j \sum_{i=1}^n r_{ij}$$

Optimizing the cost function  $\mathcal{J}$  for fixed  $r_{ik}$ 

$$\frac{\partial}{\partial \mu_j} \mathcal{J}(\mu_1, \dots, \mu_K) = -2 \sum_{i=1}^n r_{ij} \mathbf{x}_i + 2\mu_j \sum_{i=1}^n r_{ij}$$

Thus, the minimum is obtained when

$$\mu_j = \frac{\sum_i r_{ij} \mathbf{x}_i}{\sum_i r_{ij}} = \frac{1}{n_j} \sum_{\substack{i: \mathbf{x}_i \text{ is assigned} \\ \text{to cluster } j}} \mathbf{x}_i$$

where  $n_j = \sum_i r_{ij}$  is the number of points assigned to cluster j.

The optimal cluster center is given by the **centroid** (average) of all points assigned to it.

# K-means pseudocode

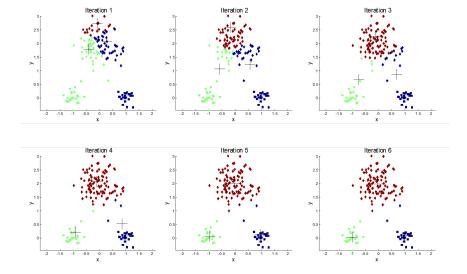
- 1. Initialize cluster centers  $\mu_1, \ldots, \mu_K$
- 2. repeat until convergence
  - ▶ assign each point to the cluster with closest center

$$r_{ik} := \begin{cases} 1 & \text{if } k = \arg\min_{k'} \|\mathbf{x}_i - \boldsymbol{\mu}_{k'}\|^2 \\ 0 & \text{otherwise} \end{cases}$$

recompute cluster centers for all k = 1, ..., K

$$\mu_k := \frac{1}{n_k} \sum_{r : r = 1} \mathbf{x}_i$$

# Example simulation K-means



## ${\bf Pre\text{-}processing:}$

- ▶ Normalize the data
- ► Remove outliers

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#### Advantages:

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#### Limitations:

- converges to local minimum . . .
- and very sensitive to initialization
  - ▶ so run it many times, and keep the **best** (i.e. sum to centers of class lower)
- sensitive to outliers and clusters of different sizes and densities
- $\triangleright$  needs number of clusters K as input
- ▶ hard cluster assignments

### K-means++

k-means++ is a variant of k-means that uses a heuristic for initializing cluster centers as follows:

- 1. Choose first center  $\mu_1$  uniformly at random from all available examples
- 2. For k = 2, ..., K
  - ▶ Choose next center  $\mu_k$  at random, where a point is chosen with probability proportional to  $\|\mathbf{x}_i \mu_l\|$ , where  $\mu_l$  is its closest cluster center picked so far (among  $\mu_1, \ldots, \mu_{k-1}$ )
- 3. Proceed with standard k-means

## Choosing the number of clusters K

The number of clusters is a hyper-parameter that has to be set by the user; unfortunately there is no obvious way to choose an  $optimal\ K$ , since oftentimes such optimal K does not exist, or is not unique, or there is no way to know.

The undelying difficulty is that in clustering there is no **true clustering** (known or unknown) so there is nothing we can compare against.

Despite this, there are many reasonable cluster quality criteria that can be used for selecting K.

These criteria measure a balance between separation of clusters and their comptactness; there is no measure that works for all datasets, and it is up to the preferences of the analyst and/or practical considerations of the problem at hand that one or another is used.

Popular ones are the Calinski-Harabasz index, the silhouette index, or the Davis-Bouldin score, but many others exist.<sup>2</sup>

<sup>&</sup>lt;sup>2</sup>This paper contains an empirical comparison of many existing indices for cluster evaluation.\

#### The Calinski-Harabasz index

The CH index uses, like k-means, Euclidean distances to measure cluster quality and so it is very much used together with k-means.

It measures the ratio between separation of cluster centers (sum of distances of cluster centers to overall mean - in the numerator) and cluster compactness (sum of distances from each point to its assigned cluster center - in the denominator):

$$\frac{(N-K)}{(K-1)} \frac{\sum_{k=1}^{K} n_k \|\mu_k - \bar{\mathbf{x}}\|^2}{\sum_{k=1}^{K} \sum_{i=1}^{n} r_{ik} \|\mathbf{x}_i - \mu_k\|^2}$$

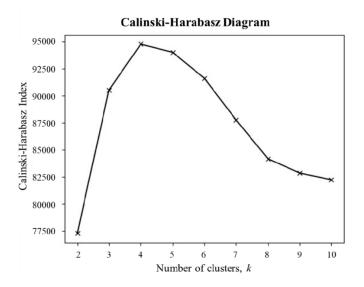
where  $\bar{\mathbf{x}}$  is the overall average of points in the dataset, i.e.  $\bar{\mathbf{x}} = \frac{1}{n} \sum_{i} \mathbf{x}_{i}$ .

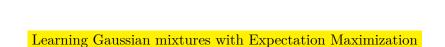
The quantities are normalized by  $\frac{(N-K)}{(K-1)}$  in order to avoid larger K having better values.

Typically, we will run k-means for different values of k, and will select the k that maximizes this index.

## The Calinski-Harabasz index

Choosing the number of clusters K

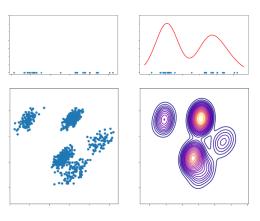




## Mixture of Gaussians

A way of modelling unknown density

When we have data that are clearly not Gaussian, it may be a useful choice to describe the data given:



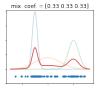
A mixture of Gaussians is a distribution that is built using a **convex sum of Gaussians**; so it is more flexible than a single Gaussian.

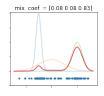
$$p(\mathbf{x}|\theta) = \sum_{k=1}^{K} \pi_k \mathcal{N}(\mathbf{x}; \mu_k, \Sigma_k)$$

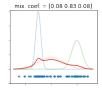
- ▶ Each  $\mathcal{N}(\mu_k, \Sigma_k)$  is a **component** of the mixture (Gaussian, with parameters  $\mu_k$  and  $\Sigma_k$ )
- ▶ The  $\pi_k$  are the **mixing coefficients** of each component, such that  $0 \le \pi_k \le 1$ , and  $\sum_k \pi_k = 1$
- The parameters of this distribution are  $\theta = \{\pi_k, \mu_k, \Sigma_k\}_{k=1,...,K}$

A key assumption of this distribution is that each data point has been generated from one component however we do not know which one.

1D Example with a mixture of 3 components, varying mixing coefficients









# Clustering with a Gaussian mixture<sup>4</sup>

One cluster == one component of mixture

So, to cluster data  $D = \{\mathbf{x}_1, \dots, \mathbf{x}_n\}$  into K clusters:

- use EM to estimate mixture, results in  $\hat{\pi}_k, \hat{\mu}_k, \hat{\Sigma}_k$  for each  $k = 1, \dots, K$
- ▶ find assignments<sup>3</sup> of each  $\mathbf{x}_i$  to clusters

 $<sup>^3</sup>$ As we will see, these assignments under the new probabilistic model are going to be **soft**. More on this when we introduce EM.

<sup>&</sup>lt;sup>4</sup>You can find a nice explanation of EM for GMM in these two blog posts part1 and part2.

Generative model perspective

To sample from such a mixture, we can think of the following **generative model**; it uses a **latent** (unobserved) variable  $\mathbf{z} = (z_1, \dots, z_K)$  whose components are all 0 except one which denotes the component from which we will sample

- 1. Pick **component** k with probability  $\pi_k$  (that is,  $z_k = 1$  w.p.  $\pi_k$ )
- 2. Generate sample **x** according to  $\mathcal{N}(\mu_k, \Sigma_k)$

The probability of generating a sample  ${\bf x}$  using this generative model is

$$p(\mathbf{x}) = \sum_{\mathbf{z}} p(\mathbf{x}, \mathbf{z}) = \sum_{k} p(\mathbf{x}, z_k = 1) = \sum_{k} p(\mathbf{x}|z_k = 1) p(z_k = 1) = \sum_{k} \pi_k \mathcal{N}(\mathbf{x}; \mu_k, \Sigma_k)$$

The joint distribution of  $\mathbf{x}$  and  $\mathbf{z}$  is thus given by (notational trick):

$$p(\mathbf{x}, \mathbf{z}) = \prod_{k} \pi_k^{z_k} \mathcal{N}(\mathbf{x}; \mu_k, \Sigma_k)^{z_k}$$

▶ joint distribution

$$p(\mathbf{x}, \mathbf{z}) = \prod_{k} \pi_k^{z_k} \mathcal{N}(\mathbf{x}; \mu_k, \Sigma_k)^{z_k}$$

marginal distribution over x

$$\begin{split} p(\mathbf{x}) &= \sum_{k} \pi_{k} \mathcal{N}(\mathbf{x}; \mu_{k}, \Sigma_{k}) \\ &= \sum_{\mathbf{z}} p(\mathbf{x}, \mathbf{z}) = \sum_{\mathbf{z}} \prod_{k'} \pi_{k'}^{z_{k'}} \mathcal{N}(\mathbf{x}; \mu_{k'}, \Sigma_{k'})^{z_{k'}} \end{split}$$

▶ marginal distribution over  $\mathbf{z}$   $p(z_k = 1) = \pi_k$  for all k = 1, ..., K and so:

$$p(\mathbf{z}) = \prod_k \pi_k^{z_k}$$

conditional distribution of x given z

$$p(\mathbf{x}|z_k = 1) = \mathcal{N}(\mathbf{x}; \mu_k, \Sigma_k)$$

Using Bayes, we compute the **conditional distribution of z given x**:

$$p(z_k = 1|\mathbf{x}) = \frac{p(\mathbf{x}|z_k = 1)p(z_k = 1)}{p(\mathbf{x})}$$

$$= \frac{p(\mathbf{x}|z_k = 1)p(z_k = 1)}{\sum_{k'} \pi_{k'} \mathcal{N}(\mathbf{x}; \mu_{k'}, \Sigma_{k'})}$$

$$= \frac{\pi_k \mathcal{N}(\mathbf{x}; \mu_k, \Sigma_k)}{\sum_{k'} \pi_{k'} \mathcal{N}(\mathbf{x}; \mu_{k'}, \Sigma_{k'})}$$

$$=: \gamma_k(\mathbf{x})$$



The quantity  $\gamma_k(\mathbf{x}) := p(z_k = 1|\mathbf{x})$  states how probable it is that a particular data point  $\mathbf{x}$  has been generated by mixture component k. Or, in the context of clustering, how probable it is that  $\mathbf{x}$  belongs to cluster k.

We use these quantities as **soft membership** to each cluster. If you want a **hard membership** then **x** should be assigned to  $k = \arg \max_{k'} \gamma_{k'}(\mathbf{x})$ . But in many contexts having soft memberships is desirable and certainly more flexible.

We are given an iid sample of unlabelled observations  $D = \{\mathbf{x}_1, \dots, \mathbf{x}_n\}$  with each  $\mathbf{x}_i \in \mathbb{R}^d$ . We want to model this sample as a Gaussian mixture. The unknown parameters are  $\theta = \{\pi_k, \mu_k, \Sigma_k\}_k$ ; K is assumed fixed and given as input.

$$l(\theta) = \log \mathcal{L}(\theta) = \log \prod_{i=1}^{K} p(\mathbf{x}; \theta)$$
$$= \log \prod_{i} \sum_{k} \pi_{k} \mathcal{N}(\mathbf{x}; \mu_{k}, \Sigma_{k})$$
$$= \sum_{i} \log \left\{ \sum_{k} \pi_{k} \mathcal{N}(\mathbf{x}; \mu_{k}, \Sigma_{k}) \right\}$$

This is hard to optimize; the log-likelihood surface is complex with many local maxima.

Maximum likelihood estimate for  $\mu_k$ 

$$l(\theta) = \sum_{i} \log \left\{ \sum_{k} \pi_{k} \mathcal{N}(\mathbf{x}; \mu_{k}, \Sigma_{k}) \right\}$$
$$= \sum_{i} \log \left\{ \sum_{\mathbf{z}} \prod_{k'} \pi_{k'}^{z_{k'}} \mathcal{N}(\mathbf{x}; \mu_{k'}, \Sigma_{k'})^{z_{k'}} \right\}$$

Nevertheless, we can compute some partial derivatives to see what conditions should hold in any local maximum.

$$ightharpoonup \frac{\partial l(\theta)}{\partial \mu_k} = 0$$
 leads to

$$\hat{\mu}_k = \frac{\sum_i \gamma_k(\mathbf{x}_i) \mathbf{x}_i}{\sum_i \gamma_k(\mathbf{x}_i)} = \frac{\sum_i P(z_k = 1 | \mathbf{x}_i) \mathbf{x}_i}{\sum_i P(z_k = 1 | \mathbf{x}_i)}$$

which is a **weighted average** of the points in our data, with weights being the **soft** assignments of each point to cluster k.

**Problem:** we do not know  $\gamma_k(\mathbf{x})$  without  $\mu_k, \Sigma_k, \pi_k$ .

Maximum likelihood estimate for  $\Sigma_k$ 

$$l(\theta) = \sum_{i} \log \left\{ \sum_{k} \pi_{k} \mathcal{N}(\mathbf{x}; \mu_{k}, \Sigma_{k}) \right\}$$

 $ightharpoonup \frac{\partial l(\theta)}{\partial \Sigma_k} = 0$  leads to

$$\hat{\Sigma}_k = \frac{\sum_i \gamma_k(\mathbf{x}_i)(\mathbf{x}_i - \hat{\mu}_k)(\mathbf{x}_i - \hat{\mu}_k)^T}{\sum_i \gamma_k(\mathbf{x}_i)} = \frac{\sum_i P(z_k = 1 | \mathbf{x}_i)(\mathbf{x}_i - \hat{\mu}_k)(\mathbf{x}_i - \hat{\mu}_k)^T}{\sum_i P(z_k = 1 | \mathbf{x}_i)}$$

which is the sample covariance matrix of all  $\mathbf{x}_i$  weighted by the soft assignments of ech point to cluster k (i.e., weighted by the posterior probability that component k generated  $\mathbf{x}_i$ )

**Problem:** we do not know  $\gamma_k(\mathbf{x})$  without  $\mu_k, \Sigma_k, \pi_k$ .

Maximum likelihood estimate for  $\pi_k$ 

Now we maximize the **Lagrangian**  $l(\theta) - \lambda(\sum_k \pi_k - 1)$  since we have an equality constraint on the  $\pi_k$ 

▶  $\frac{\partial l(\theta)}{\partial \pi_k} = 0$  and  $\sum_k \pi_k = 1$  lead to

$$\hat{\pi}_k = \frac{1}{n} \sum_{i} \gamma_k(\mathbf{x}_i)$$

which is the average of all soft assignments for all data point  $x_i$ .

**Problem:** we do not know  $\gamma_k(\mathbf{x})$  without  $\mu_k, \Sigma_k, \pi_k$ .

- $\blacktriangleright$  We can estimate  $\pi_k, \Sigma_k, \mu_k$  if we know  $\gamma_k(\cdot)$
- ▶ We can compute  $\gamma_k(\cdot)$  from estimates  $\hat{\pi}_k, \hat{\Sigma}_k, \hat{\mu}_k$

# Learning Gaussian mixtures with EM, cont. Pseudocode

- 1. Initialize  $\hat{\mu}_k, \hat{\Sigma}_k, \hat{\pi}_k$
- 2. repeat until convergence
  - **E-step** recompute soft assignments  $\gamma_k(\mathbf{x}_i)$

$$\gamma_k(\mathbf{x}_i) = \frac{\pi_k \mathcal{N}(\mathbf{x}_i; \mu_k, \Sigma_k)}{\sum_{k'} \pi_{k'} \mathcal{N}(\mathbf{x}_i; \mu_{k'}, \Sigma_{k'})}$$

▶ M-step recompute ML estimates

$$\hat{\mu}_k = \frac{\sum_i \gamma_k(\mathbf{x}_i) \mathbf{x}_i}{\sum_i \gamma_k(\mathbf{x}_i)}$$

$$\hat{\Sigma}_k = \frac{\sum_i \gamma_k(\mathbf{x}_i) (\mathbf{x}_i - \hat{\mu}_k) (\mathbf{x}_i - \hat{\mu}_k)^T}{\sum_i \gamma_k(\mathbf{x}_i)}$$

$$\hat{\pi}_k = \frac{1}{n} \sum_i \gamma_k(\mathbf{x}_i)$$

Learning Gaussian mixtures with EM, cont. Initialization

Commonly we initialize  $\hat{\mu}_k, \hat{\Sigma}_k, \hat{\pi}_k$  using the result of k-means:

- 1. Run k-means with k=K (maybe a few times, pick best)
- 2. Set  $\hat{\mu}_k$  to k-mean's cluster centers
- 3. Set each  $\hat{\Sigma}_k$  to the sample covariance of each cluster of k-means
- 4. Set  $\hat{\pi}_k$  as the fraction of examples assigned to cluster k

#### Model selection

#### Selection of K in mixture models:

- $\blacktriangleright$  Cannot just choose K that maximizes likelihood because  $l(\theta)$  is always larger for larger K
- Penalize complexity
  - $\blacktriangleright$  f.i. Bayesian information criterion (BIC):  $BIC = l(\theta) d/2\log n$  , where d is # parameters
  - Asymptotically correct under certain assumptions
  - Often used in practice for mixture models even though assumptions for theory are not met

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#### Model selection with same K:

- ► Cross validation:
  - Score different models by  $\log p(Xtest|\theta)$
  - split data into train and validate sets
  - Works well on large data sets
  - Can be noisy on small data (log L is sensitive to outliers)

# EM, special cases

We can restrict the **shape** of Gaussians for each component, which results in special cases of mixtures:

- No restriction on  $\Sigma_k$ ; this is the general case (most flexible); each cluster can have general Gaussian shape
- $\Sigma_k$  are diagonal; each Gaussian component is forced to have no correlation among input dimensions (i.e. axis-aligned)
- $\Sigma_k = \sigma^2 I$  are **isotropic** or **spherical**; each Gaussian component is forced to be spherical, so no correlation among input variables and same scaling accross each input variable.

In fact, k-means is a degenerate case of this scenario: if  $\sigma^2 \to 0$ , then  $\gamma_k(\mathbf{x}_i) \to r_{ik}$