#### The degree distribution

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#### Visual fitting

Non-linear regression

Likelihood

The challenge of parsimony

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## The limits of visual analysis

A syntactic dependency network [Ferrer-i-Cancho et al., 2004]



## The empirical degree distribution

- ► *N*: finite number of vertices / *k* vertex degree
- n(k): number of vertices of degree k.
- ► n(1),n(2),...,n(N) defines the degree spectrum (loops are allowed).
- ► n(k)/N: the proportion of vertices of degree k, which defines the (empirical) degree distribution.
- ▶ p(k): function giving the probability that a vertex has degree k, p(k) ≈ n(k)/N.

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• p(k): probability mass function (pmf).

#### Example: degree spectrum



- Global syntactic dependency network (English)
- Nodes: words
- Links: syntactic dependencies

#### Not as simple:

- Many degrees occurring just once!
- Initial bending or hump: power-law? ▲ □ ► < □ ►</p>

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#### Example: empirical degree distribution



- Notice the scale of the y-axis.
- Normalized version of the degree spectrum (dividing over N).

# Example: in-degree (red) degree versus out-degree (green)



- The distribution of in-degree and that of out-degree do not need to be identical!
- Similar for global syntactic dependency networks?
   Differences in the distribution or the parameters?
- Known cases of radical differences between in and out-degree distributions (e.g., web pages, wikipedia articles).
   In-degree more power-law like than out-degree.

# What is the mathematical form of p(k)?

Possible degree distributions

- ► The typical hypothesis: a power-law p(k) = ck<sup>-γ</sup> but what exactly? How many free parameters?
  - Zeta distribution: 1 free parameter.
  - ▶ Right-truncated zeta distribution: 2 free parameters.

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Motivation:

- Accurate data description (looks are deceiving).
- Help to design or select dynamical models.

# Zeta distributions I

Zeta distribution:

$$p(k) = \frac{1}{\zeta(\gamma)}k^{-\gamma},$$

being

$$\zeta(\gamma) = \sum_{x=1}^{\infty} x^{-\gamma}$$

the Riemann zeta function.

 (here it is assumed that γ is real) ζ(γ) converges only for γ > 1 (γ > 1 is needed).

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- $\gamma$  is the only free parameter!
- Do we wish p(k) > 0 for k > N?

# Zeta distributions I

Right-truncated zeta distribution

$$p(k) = \frac{1}{H(k_{\max},\gamma)}k^{-\gamma},$$

being

$$H(k_{max},\gamma) = \sum_{x=1}^{k_{max}} x^{-\gamma}$$

the generalized harmonic number of order  $k_{max}$  of  $\gamma$ . Or why not

$$p(k) = ck^{-\gamma}e^{-keta}$$

(modified power-law, Altmann distribution,...) with 2 or 3 free parameters? Which one is best? (standard model selection)

# What is the mathematical form of p(k)?

Possible degree distributions

The null hypothesis (for a Erdös-Rényi graph without loops)

$$p(k) = \binom{N-1}{k} \pi^k (1-\pi)^{N-1-k}$$

with  $\pi$  as the only free parameter (assuming that N is given by the real network).

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Binomial distribution with parameters N - 1 and  $\pi$ , thus  $\langle k \rangle = (N - 1)\pi \approx N\pi$ .

► Another null hypothesis: random pairing of vertices with constant number of edges *E*.

# The problems II

- Is f(k), a good candidate? Does f(k) fit the empirical degree distribution well enough?
- ► *f*(*k*) is a (candidate) model.
- How do we evaluate goodness of a model? Three major approaches:
  - Qualitatively (visually).
  - The error of the model: the deviation between the model and the data.
  - The likelihood of the model: the probability that the model produces the data.

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# Visual fitting

Assume a two variables: a predictor x (e.g., k, vertex degree) and a response y (e.g., n(k), the number vertices of degree k; or p(k)...).

- Look for a transformation of the at least one of the variables showing approximately a straight line (upon visual inspection) and obtain the dependency between the two original variables.
- Typical transformations: x' = log(x), y' = log(y).

1. If 
$$y' = log(y) = ax + b$$
 (linear-log scale) then  
 $y = e^{ax+b} = ce^{ax}$ , with  $c = e^b$  (exponential).

2. If 
$$y' = log(y) = ax' + b = alog(x) + b$$
 (log-log scale) then  $y = e^{alog(x)+b} = cx^a$ , with  $c = e^b$  (power-law).

3. If y = ax' + b = alog(x) + b (log-linear scale) then the transformation is exactly the functional dependency between the original variables (logarithmic).

#### What is this distribution?



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## Solution: geometric distribution

 $y = (1 - p)^{x-1}p$  (with p = 1/2 in this case). In standard exponential form,

$$y = (1-p)^{x} \frac{p}{1-p} = e^{x \log(1-p)} \frac{p}{1-p} = ce^{ax}$$

with  $a = \log(1 - p)$  and c = p/(1 - p). Examples:

- Random network models (degree is geometrically distributed).
- Distribution of word lengths in random typing (empty words are not allowed) [Miller, 1957].
- Distribution of projection lengths in real neural networks [Ercsey-Ravasz et al., 2013].

#### A power-law distribution



What is the exponent of the power-law?

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#### Solution: zeta distribution

$$y = \frac{1}{\zeta(a)} x^{-a}$$

with a = 2. Formula for  $\zeta(a)$  is known for certain integer values, e.g.,  $\zeta(2) = \pi^2/6 \approx 1.645$ . Examples:

- Empirical degree distribution of global syntactic dependency networks [Ferrer-i-Cancho et al., 2004] (but see also lab session on degree distributions).
- ▶ Frequency spectrum of words in texts [Corral et al., 2015].

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#### What is this distribution?



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#### Solution: a "logarithmic" distribution

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$$y = c(log(x_{max}) - \log x))$$

with  $x = 1, 2, ..., x_{max}$  and c being a normalization term, i.e.

$$c = \frac{1}{\sum_{x=1}^{x_{max}} \left( \log(x_{max}) - \log x) \right)}$$

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## The problems of visual fitting

- The right transformation to show linearity might not be obvious (taking logs is just one possibility).
- Looks can be deceiving with noisy data.
- A good guess or strong support for the hypothesis requires various decades.

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Solution: a quantitative approach.

# Non-linear regression I [Ritz and Streibig, 2008]

- A univariate response y.
- A predictor variable x
- ► Goal: functional dependency between *y* and *x*.

Formally:  $y = f(x, \beta)$ , where

- $f(x,\beta)$  is the "model".
- K parameters.
- $\flat \ \beta = (\beta_1, ..., \beta_K)$

Examples:

- Linear model: f(x, (a, b)) = ax + b (K = 2).
- A non-linear model (power-law):  $f(x, (a, b)) = ax^b$  (K = 2).

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# Non-linear regression II

Problem of regression:

- ► A data set of n pairs: (x<sub>1</sub>, y<sub>1</sub>), ..., (x<sub>n</sub>, y<sub>n</sub>). Example: x<sub>i</sub> is vertex degree (k) and y<sub>i</sub> is the number of vertices of degree k (n(k)) of a real network.
- n is the sample size.
- ► f(x, β) is unlikely to give a perfect fit. y<sub>1</sub>, y<sub>2</sub>, ..., y<sub>n</sub> may contain error.

Solution: the conditional mean response

$$\mathsf{E}(y_i|x_i)=f(x_i,\beta)$$

 $(f(x,\beta)$  is not actually the model for the data points but a model for expectation given  $x_i$ ).

## Non-linear regression II

The full model is then

$$y_i = E(y_i|x_i) + \epsilon_i = f(x_i, \beta) + \epsilon_i$$

The quality of the fit of a model with certain parameters: the residual sums of squares

$$RSS(\beta) = \sum_{i=1}^{n} (y_i - f(x_i, \beta))^2$$

The parameters of the model are estimated minimizing the RSS. Non-linear regression: minimization of RSS.

Common metric of the quality of the fit: the residual standard error

$$s^2 = \frac{RSS(\beta)}{n-K}$$

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#### Example of non-linear regression



- Non-linear regression yields y = 2273.8x<sup>-1.23</sup> (is the exponent that low?)
- Is the method robust? (=not distracted by undersampling, noise, and so on)
- Likely and unlikely events are weighted equally.
- Solution: weighted regression, taking likelihood into account,...

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# Likelihood I [Burnham and Anderson, 2002]

- A probabilistic metric of the quality of the fit.
- L(parameters|data, model): likelihood of the parameters given the data (sample of size n) and a model.
   Example: L(γ|data, Zeta distribution with parameterγ)

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 Best parameters: the parameters that maximize L(parameters|data, model).

# Likelihood II

- Consider a sample x<sub>1</sub>, x<sub>2</sub>, ...x<sub>n</sub> (e.g., the degree sequence of a network).
- Definition (assuming independence)

$$L(parameters|data, model) = \prod_{i=1} p(x_i; parameters)$$

For a zeta distribution

$$L(\gamma | x_1, x_2, ..., x_n; \text{Zeta distribution}) = \prod_{i=1}^n p(x_i; \gamma)$$
$$= \zeta(\gamma)^{-n} \prod_{i=1}^n x_i^{-\gamma}$$

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## Log-likelihood

Likelihood is a vanishingly small number. Solution: taking logs.

 $\mathcal{L}(parameters | data, model) = \log L(parameters | data, model) \\ = \sum_{i=1} \log p(x_i; parameters)$ 

Example:

$$\mathcal{L}(\gamma | x_1, x_2, ..., x_n; \text{Zeta distribution}) = \sum_{i=1}^n \log p(x_i; \gamma)$$
$$= \gamma \sum_{i=1}^n \log x_i - n \log(\zeta(\gamma))$$

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Question to the audience

# What is the best model for data?

Cue: a universal method.

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# What is the best model for data?

- The best model of the data is the data itself. Overfitting!
- The quality of the fit cannot decrease if more parameters are added (wisely). Indeed, the quality of the fit normally increases when adding parameters.
- The metaphor of picture compression. Compressing a picture (with quality reduction). A good compression technique shows a nice trade-off between file size and image quality).
- Modelling is compressing a sample, the empirical distribution (e.g., compressing the degree sequence of a network).
  - Models with many parameters should be penalized!
  - Models compressing the data with a low quality should be also penalized.

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How?

# Akaike's information criterion (AIC)

$$AIC = -2\mathcal{L} + 2K,$$

with K being the number of parameters of the model. For small samples, a correction is necessary

$$AIC_c = -2\mathcal{L} + 2\mathcal{K}\left(\frac{n}{n-\mathcal{K}-1}\right),$$

or equivalently

$$AIC_{c} = -2\mathcal{L} + 2K + \frac{2K(K+1)}{n-K-1}$$
$$= AIC + \left(\frac{2K(K+1)}{n-K-1}\right)$$

 $AIC_{c}$  is recommended if  $n \gg K$  is not satisfied!

#### Model selection with AIC

- What is the best of a set of models? The model that minimizes AIC
- ► *AIC*<sub>best</sub>: the AIC of the model with smallest *AIC*.
- Δ: "AIC difference", the difference between the AIC of the model and that of the best model (Δ = 0 for the best model).

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#### Example of model selection with AIC

Consider the case of model selection with three nested models: Model 1  $p(k) = \frac{k^{-2}}{\zeta(2)}$  (zeta distribution with (-)2 exponent) Model 2  $p(k) = \frac{k^{-\gamma}}{\zeta(\gamma)}$  (zeta distribution) Model 3  $p(k) = \frac{k^{-\gamma}}{H(k_{max},\gamma)}$  (right-truncated zeta distribution)

Model *i* is nested model of i - 1 if the model *i* is a generalization of model i - 1 (adding at least one parameter).

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#### Example of model selection with AIC

Model	Κ	$\mathcal{L}$	AIC	Δ
1	0			
2	1			
3	2			

Imagine that the true model is a zeta distribution with  $\gamma=1.5$  and the sample is large enough, then

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Model	Κ	$\mathcal{L}$	AIC	Δ
1	0			≫ 0
2	1			0
3	2			> 0

# AIC for non-linear regression I

- RSS: "distance" between the data and fitted regression curve based on the the model fit.
- AIC: estimate of the "distance" from the model fit to the true but unknown model that generated the data.
- In a regression model one assumes that the error e follows a normal distribution, the p.d.f. is

$$f(\epsilon) = \frac{1}{(2\pi\sigma^2)^{1/2}} exp\left\{-\frac{(\epsilon-\mu)^2}{2\sigma^2}\right\}$$

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The only parameter is  $\sigma$  as standard non-linear regression assumes  $\mu = 0$ .

#### AIC for non-linear regression II

• Applying 
$$\mu = 0$$
 and  $\epsilon_i = y_i - f(x_i, \beta)$ 

$$f(\epsilon_i) = \frac{1}{(2\pi\sigma^2)^{1/2}} \exp\left\{-\frac{(y_i - f(x_i,\beta))^2}{2\sigma^2}\right\}$$

Likelihood in a regression model:

$$L(\beta,\sigma^2) = \prod_{i=1}^n f(\epsilon_i)$$

After some algebra one gets

$$L(\beta,\sigma^2) = \frac{1}{(2\pi\sigma^2)^{n/2}} exp\left\{-\frac{RSS(\beta)}{2\sigma^2}\right\}$$

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# AIC for non-linear regression III

Equivalence between maximization of likelihood and minimization of error (under certain assumptions)

• If  $\hat{\beta}$  is the **best estimate** of  $\beta$  then

$$L(\hat{\beta}, \hat{\sigma}^2) = \frac{1}{(2\pi RSS(\hat{\beta})/n)^{n/2}} exp(-n/2)$$

thanks to  $\hat{\sigma}^2 = \frac{n-K}{n}s^2$  (recall  $s^2 = \frac{RSS(\beta)}{n-K}$ ).

Models selection with regression models:

$$AIC = -2 \log L(\hat{\beta}, \hat{\sigma}^2)) + 2(K+1) = n \log(2\pi) + n \log(RSS(\hat{\beta})/n) + n + 2(K+1)$$

Why the term for parsimony is 2(K + 1) and not K?

# Concluding remarks

- Under non-linear regression AIC is the way to go for model selection if the models are not nested (alternative methods do exist for nested models [Ritz and Streibig, 2008]).
- Equivalence between maximum likelihood and non-linear regression implies some assumption (e.g., homocedasticity).

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