

Generating Random Derangements

Conrado Martínez¹ Alois Panholzer² Helmut Prodinger³

¹Univ. Politècnica de Catalunya, Spain

²Tech. Univ. Wien, Austria

³Univ. Stellenbosch, South Africa

Derangements

Le Problème des Derangements:

“A number of gentlemen, say n , surrender their top hats in the cloakroom and proceed to the evening's enjoyment. After wining and dining (and wining some more), they stumble back to the cloakroom and confusedly take the first top-hat they see. What is the probability that no gentleman gets his own hat?”

Derangements

- ▶ A **derangement** is a permutation without fixed points: $\pi(i) \neq i$ for any i , $1 \leq i \leq n$
- ▶ The number D_n of derangements of size n is

$$D_n = n! \cdot \left[\frac{1}{0!} - \frac{1}{1!} + \frac{1}{2!} - \frac{1}{3!} + \cdots + \frac{(-1)^n}{n!} \right] = \left\lfloor \frac{n! + 1}{e} \right\rfloor.$$

- ▶ As $n \rightarrow \infty$, $D_n/n! \sim 1/e \approx 0.36788$.

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- ▶ As $n \rightarrow \infty$, $D_n/n! \sim 1/e \approx 0.36788$.

Break #1: Fisher-Yates' shuffle

```
procedure RandomPermutation( $n$ )
    for  $i \leftarrow 1$  to  $n$  do  $A[i] \leftarrow i$ 
    for  $i \leftarrow n$  downto 1 do
         $j \leftarrow \text{Uniform}(1, i)$ 
         $A[i] \leftrightarrow A[j]$ 
    return  $A$ 
```

Break #2: Sattolo's algorithm

```
procedure RandomCyclicPermutation( $n$ )
    for  $i \leftarrow 1$  to  $n$  do  $A[i] \leftarrow i$ 
    for  $i \leftarrow n$  downto 1 do
         $j \leftarrow \text{Uniform}(1, i - 1)$ 
         $A[i] \leftrightarrow A[j]$ 
    return  $A$ 
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A recurrence for the number of derangements

$$D_0 = 1, D_1 = 0$$

$$D_n = (n - 1)D_{n-1} + (n - 1)D_{n-2}$$

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Choice #1: n belongs to a cycle of length > 2 .

The derangement of size n is built by constructing a derangement of size $n - 1$ and then n is inserted into any of the cycles (of length ≥ 2); there are $(n - 1)$ possible ways to do that

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Choice #2: n belongs to a cycle of length 2.

The derangement of size n is built by constructing a cycle of size 2 with n and some j , $1 \leq j \leq n - 1$; then we build a derangement of size $n - 2$ with the remaining elements

The rejection method

Require: $n \neq 1$

```
procedure RandomDerangement( $n$ )
    repeat
         $A \leftarrow \text{RandomPermutation}(n)$ 
    until Is-Derangement( $A$ ) return  $A$ 
```

$$\mathbb{P}[A \text{ is a derangement}] \approx \frac{1}{e}$$

$$\mathbb{E}[\#\text{ of calls to Random}] = e \cdot n + O(1)$$

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The recursive method

$$C \leftarrow \{1, 2, \dots, n\}$$

RandomDerangement-Rec(n, C)

Require: $n \neq 1$

procedure RandomDerangement-Rec(n, C)

if $n \leq 1$ **then return**

$j \leftarrow$ a random element from C

$p \leftarrow$ Uniform(0, 1)

if $p < (n - 1)D_{n-2}/D_n$ **then**

 RandomDerangement-Rec($n - 2, C \setminus \{j, n\}$)

$\pi(n) \leftarrow j; \pi(j) \leftarrow n$

else

 RandomDerangement-Rec($n - 1, C \setminus \{n\}$)

$\pi(n) \leftarrow \pi(j); \pi(j) \leftarrow n$

Our algorithm

Require: $n \neq 1$

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procedure RandomDerangement( $n$ )
    for  $i \leftarrow 1$  to  $n$  do  $A[i] \leftarrow i$ ;  $mark[i] \leftarrow \text{false}$ 
     $i \leftarrow n$ ;  $u \leftarrow n$ 
    while  $u \geq 2$  do
        if  $\neg mark[i]$  then
             $j \leftarrow$  pick a random unmarked element in  $A[1..i - 1]$ 
             $A[i] \leftrightarrow A[j]$ 
            if  $j$  shall be marked  $\equiv$  close the cycle then
                 $mark[j] \leftarrow \text{true}$ ;  $u \leftarrow u - 1$ 
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            repeat  $j \leftarrow \text{Random}(1, i - 1)$ 
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The analysis

- ▶ # of marked elements = # of cycles (C_n)
- ▶ # of iterations = # of calls to Uniform = $n - C_n$
- ▶ G = # of calls to Random
- ▶ G_i = # of calls to Random at iteration i
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$$\begin{aligned}\mathbb{E}[\text{cost}] &= n - \mathbb{E}[C_n] + \mathbb{E}[G] \\ &= n - \mathbb{E}[C_n] + \sum_{1 \leq i \leq n} \mathbb{E}[G_i]\end{aligned}$$

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The computation of $\mathbb{E}[C_n]$ can be done via standard generating function techniques:

$$\begin{aligned} C(z, v) &= \sum_{A \in \mathcal{D}} \frac{z^{|A|}}{|A|!} v^{\# \text{ cycles}(A)} \\ &= \exp \left(v \left(\log \frac{1}{1-z} - z \right) \right) = e^{-vz} \frac{1}{(1-z)^v} \\ \mathbb{E}[v^{C_n}] &= \frac{n!}{D_n} [z^n] C(z, v) = \frac{e^{1-v}}{(v-1)!} n^{v-1} (1 + O(n^{-1+\epsilon})) \\ \mathbb{E}[C_n] &= \log n + O(1), \quad \mathbb{V}[C_n] = \log n + O(1) \\ \frac{C_n - \log n}{\sqrt{\log n}} &\rightarrow \mathcal{N}(0, 1) \end{aligned}$$

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The analysis

- ▶ M_i indicator variable for the event “ $A[i]$ gets marked”
- ▶ $M_i = 1 \implies G_i = 0$
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$$\mathbb{E}[G] = \sum_{1 < i \leq n} \mathbb{E}[G_i | M_i = 0] \cdot \mathbb{P}[M_i = 0]$$

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- ▶ $U_i = \#$ of unmarked elements in $A[1..i]$; $U_n = n$
- ▶ $B_{i+1} = \#$ of marked elements in $A[1..i]$; $B_{n+1} = 0$
- ▶ $U_i + B_{i+1} = i$
- ▶ If $A[i]$ is not marked then G_i is geometrically distributed with probability of success
 $(U_i - 1)/(i - 1) = (i - 1 - B_{i+1})/(i - 1)$; hence

$$\mathbb{E}[G_i \mid M_i = 0] = \mathbb{E}\left[\frac{i - 1}{i - 1 - B_{i+1}} \mid M_i = 0\right]$$

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- ▶ $B_{i+1} \leq C_n$
- ▶ $0 \leq B_{i+1} \leq i$
- ▶ $U_i \neq 1$ and $B_{i+1} \neq 1$ for all $1 \leq i \leq n$
- ▶ If $M_i = 0$ then $B_{i+1} < i - 1$

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The analysis

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$$\begin{aligned}\mathbb{E}[G] &= \sum_{1 < i \leq n} \mathbb{E}\left[\frac{i-1}{i-1-B_{i+1}} \mid M_i = 0\right] \cdot \mathbb{P}[M_i = 0] \\ &\leq \sum_{1 < i \leq n} \mathbb{E}\left[\min\left\{i-1, \frac{i-1}{i-1-C_n}\right\}\right] \\ &\leq \sum_{1 \leq k \leq \lfloor n/2 \rfloor} \mathbb{P}[C_n = k] \left(\sum_{i=1}^{k+1} (i-1) + \sum_{i=k+2}^{\lfloor n/2 \rfloor} \frac{i-1}{i-1-k} \right) \\ &= n - 1 - \mathbb{E}[C_n] + \frac{1}{2} \mathbb{E}[C_n^2] + O(\mathbb{E}[C_n \log(n - C_n)]) \\ &= n + O(\mathbb{E}[C_n^2]) + O(\log n \cdot \mathbb{E}[C_n]) = n + O(\log^2 n)\end{aligned}$$

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The analysis

Since we also have $\mathbb{E}[G] \geq n - \mathbb{E}[C_n]$, we have finally

$$\mathbb{E}[\text{cost}] = n - \mathbb{E}[C_n] + \mathbb{E}[G] = 2n + O(\log^2 n)$$